

Time Series Analysis

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Class 11

Forecasting with ARMA

- After the model is chosen and diagnostic is done, the $ARMA(p, q)$ can be used to forecast future values of the time series.
- In probabilistic terms, the forecast is considered on the basis of the collected information on the process at time t , denoted \mathcal{F}_t , where

$$\mathcal{F}_t = \{X_t, X_{t-1}, X_{t-2}, \dots, X_0, X_{-1}, \dots\}.$$

- **Goal:** To forecast the value of the process at time $t + k$, X_{t+k} , conditioning on \mathcal{F}_t .
- Define \hat{X}_{t+k} the predictor of X_{t+k} and \hat{x}_{t+k} its realization.
- Choose \hat{X}_{t+k} such that it **minimizes the forecast men squared error**

$$\mathbb{E}(e_{t+k}^2) = \mathbb{E}(X_{t+k} - \hat{X}_{t+k})^2,$$

where $e_{t+k} = X_{t+k} - \hat{X}_{t+k}$ is the **forecast error** and is a measure of the model predictive ability.

- It can be shown that the **best** predictor \hat{X}_{t+k} is given by its conditional expected value:

$$\hat{X}_{t+k} = \mathbb{E}(X_{t+k}|\mathcal{F}_t) = \mathbb{E}_t(X_{t+k}).$$

- Consider that for conditional expected values we have

$$\mathbb{E}(X_{t+j}|\mathcal{F}_t) = x_{t+j} \quad \text{if } j \leq 0$$

$$\mathbb{E}(X_{t+j}|\mathcal{F}_t) = \hat{x}_{t+j} \quad \text{if } j > 0$$

- Setting $e_t = X_t - \hat{X}_t$,

$$\mathbb{E}(\epsilon_{t+j}|\mathcal{F}_t) = e_{t+j} \quad \text{if } j \leq 0$$

$$\mathbb{E}(\epsilon_{t+j}|\mathcal{F}_t) = 0 \quad \text{if } j > 0$$

- Recall that e_{t+k} is a random variable thus, it is worth to evaluate its probabilistic structure.

Parameter estimation for ARMA processes

- For the ARMA process it can be shown that as the forecasting horizon increases ($k \rightarrow \infty$):
- the forecast tends to the unconditional mean of the process

$$\lim_{k \rightarrow \infty} \hat{X}_{t+j} = \mathbb{E}(X_{t+j}),$$

- the variance of the error tends to the unconditional variance of the process

$$\lim_{k \rightarrow \infty} \text{Var}(e_{t+j}) = \text{Var}(X_t).$$

- Consider a stationary $AR(1)$ process

$$X_t = \delta + \varphi X_{t-1} + \epsilon_t.$$

- If the parameters were known, the one step ahead ($k = 1$) forecast would be

$$\hat{x}_{t+1} = \mathbb{E}_t(X_{t+1}) = \mathbb{E}_t(\delta + \varphi X_t + \epsilon_{t+1}) = \delta + \varphi X_t.$$

- The forecasting error

$$e_{t+1} = X_{t+1} - \hat{X}_{t+1} = \delta + \varphi X_t + \epsilon_{t+1} - (\delta + \varphi X_t) = \epsilon_{t+1}$$

$$\Rightarrow \text{Var}(e_{t+1}) = \text{Var}(\epsilon_{t+1}) = \sigma^2.$$

- In order to forecast $k = 2$ steps ahead:

$$\begin{aligned}\hat{x}_{t+2} &= \mathbb{E}_t(X_{t+2}) = \mathbb{E}_t(\delta + \varphi X_{t+1} + \epsilon_{t+2}) = \delta + \mathbb{E}_t(\varphi X_{t+1}) \\ &= \delta + \varphi \hat{x}_{t+1} = \delta + \varphi(\delta + \varphi x_t) = \delta + \delta\varphi + \varphi^2 x_t = \delta(1 + \varphi) + \varphi^2 x_t.\end{aligned}$$

- The forecasting error becomes

$$\begin{aligned}e_{t+2} &= X_{t+2} - \hat{X}_{t+2} = \delta + \varphi X_{t+1} + \epsilon_{t+2} - \delta - \varphi \hat{X}_{t+1} \\ &= \epsilon_{t+2} + \varphi(X_{t+1} - \hat{X}_{t+1}) = \epsilon_{t+2} + \varphi e_{t+1} = \epsilon_{t+2} + \varphi \epsilon_{t+1}.\end{aligned}$$

- Therefore,

$$\text{Var}(e_{t+2}) = \text{Var}(\epsilon_{t+2} + \varphi \epsilon_{t+1}) = \sigma^2(1 + \varphi^2).$$

- By iterating the procedure, the general k -th steps ahead forecast will be

$$\hat{x}_{t+k} = \delta(1 + \varphi + \varphi^2 + \dots + \varphi^{k-1}) + \varphi^k x_t$$

$$\text{Var}(e_{t+k}) = \sigma^2(1 + \varphi^2 + \varphi^4 + \dots + \varphi^{2(k-1)}).$$

- As the horizon increases:

$$\lim_{k \rightarrow \infty} \varphi^k = 0$$

$$\lim_{k \rightarrow \infty} (1 + \varphi + \varphi^2 + \dots + \varphi^{k-1}) = \frac{1}{1 - \varphi}$$

$$\lim_{k \rightarrow \infty} (1 + \varphi^2 + \dots + \varphi^{2(k-1)}) = \frac{1}{1 - \varphi^2}$$

$$\lim_{k \rightarrow \infty} \hat{x}_{t+k} = \frac{\delta}{(1 - \varphi)}.$$

- That is, for the $AR(1)$, the **point forecast converges to the unconditional mean**. If $\delta = 0$, then the forecast converges to zero.
- The **variance tends to the unconditional variance of the process**

$$\lim_{k \rightarrow \infty} \text{Var}(e_{t+k}) = \frac{\sigma^2}{1 - \varphi^2}.$$

- The **parameters in the model are replaced by their estimates**
 $\hat{\delta}$, $\hat{\varphi}$, $\hat{\sigma}^2$.
- Same conditions hold for the $AR(p)$. (As k increases, the estimate becomes the unconditional mean of the process and the variance becomes the unconditional variance of the process.)

- Consider the $MA(1)$ process

$$X_t = \delta + \theta\epsilon_{t-1} + \epsilon_t.$$

- The one-step ahead forecast is given by

$$\begin{aligned}\hat{X}_{t+1} &= \mathbb{E}_t(X_{t+1}) = \mathbb{E}_t(\delta + \theta\epsilon_t + \epsilon_{t+1}) = \\ &= \delta + \theta\mathbb{E}_t(\epsilon_t) + \mathbb{E}_t(\epsilon_{t+1}) = \delta + \theta e_t.\end{aligned}$$

- Assuming $\epsilon_t = e_t$, the one-step ahead forecast error is

$$e_{t+1} = X_{t+1} - \hat{X}_{t+1} = \delta + \theta\epsilon_t + \epsilon_{t+1} - (\delta + \theta e_t) = \epsilon_{t+1}.$$

- The expectation is null and the variance

$$\text{Var}(e_{t+1}) = \text{Var}(\epsilon_{t+1}) = \sigma^2.$$

- The two steps ahead forecast becomes

$$\begin{aligned}\hat{x}_{t+2} &= \mathbb{E}_t(X_{t+2}) = \mathbb{E}_t(\delta + \theta\epsilon_{t+1} + \epsilon_{t+2}) = \\ &= \delta + \theta\mathbb{E}_t(\epsilon_{t+1}) + \mathbb{E}_t(\epsilon_{t+2}) = \delta.\end{aligned}$$

- The forecast with an $MA(1)$ remains constant and equal to the unconditional expectation of X_t after two steps ahead.
- The same holds for an $MA(q)$.
- In general: not trivial forecasts are obtained only for $k \leq q$. For horizon $k > q$ the forecast coincides with the unconditional mean of the model.

- The forecasting error is

$$e_{t+2} = X_{t+2} - \hat{X}_{t+2} = \delta + \theta\epsilon_{t+1} + \epsilon_{t+2} - \delta = \theta\epsilon_{t+1} + \epsilon_{t+2}$$

- It follows that

$$\text{Var}(e_{t+2}) = \text{Var}(\theta\epsilon_{t+1} + \epsilon_{t+2}) = \sigma^2(1 + \theta^2)$$

- for any step ahead with $k > 2$, the variance of the error becomes $\sigma^2(1 + \theta^2)$, that is the unconditional variance of the $MA(1)$ process.
- The same can be generalized for an $MA(q)$ process with variance $\sigma^2(1 + \theta_1^2 + \dots + \theta_q^2)$.

Forecast with ARMA(1,1)

- Consider the ARMA(1,1)

$$X_t = \delta + \varphi X_{t-1} + \theta \epsilon_{t-1} + \epsilon_t.$$

- The uno step ahead forecast is given by

$$\hat{x}_{t+1} = \mathbb{E}_t(X_{t+1}) = \mathbb{E}_t(\delta + \varphi X_t + \theta \epsilon_t + \epsilon_{t+1}) = \delta + \varphi x_t + \theta e_t.$$

- The forecast error is

$$e_{t+1} = X_{t+1} - \hat{X}_{t+1} = \epsilon_{t+1}.$$

- Its variance is

$$\text{Var}(e_{t+1}) = \sigma^2.$$

- The forecast for $k = 2$ is

$$\begin{aligned}\hat{x}_{t+2} &= \mathbb{E}_t(X_{t+2}) = \mathbb{E}_t(\delta + \varphi X_{t+1} + \theta \epsilon_{t+1} + \epsilon_{t+2}) = \delta + \varphi \hat{x}_{t+1} = \\ &= \delta(1 + \varphi) + \varphi^2 x_t + \theta \varphi e_t.\end{aligned}$$

- The forecast error is

$$\begin{aligned}e_{t+2} &= X_{t+2} - \hat{X}_{t+2} = \delta + \varphi X_{t+1} + \theta \epsilon_{t+1} + \epsilon_{t+2} - (\delta + \varphi \hat{X}_{t+1}) = \\ &= \varphi e_{t+1} + \theta \epsilon_{t+1} + \epsilon_{t+2} = (\varphi + \delta) \epsilon_{t+1} + \epsilon_{t+2}.\end{aligned}$$

- Notice that it has null expected value and variance equal to

$$\text{Var}(e_{t+2}) = \sigma^2((\varphi + \delta)^2 + 1).$$

- By iterating, the k steps ahead forecast is given by

$$\hat{x}_{t+k} = \delta(1 + \varphi + \varphi^2 + \dots + \varphi^{k-1}) + \varphi^k x_t + \theta \varphi^{k-1} e_t.$$

- As k goes to ∞ the variance tends to the unconditional variance of the process.
- Notice that after the second step ahead the predictor resembles that of an $AR(1)$. Indeed, its asymptotic behaviour is exactly that of an $AR(1)$.



- When $k > 1$ the behaviour of the forecast is dominated by the autoregressive part.

- For a general $ARMA(p, q)$ similar results to those seen for the $ARMA(1, 1)$ can be obtained.
- When $k > q$ the autoregressive part drives the forecast that converges to the unconditional mean of the $ARMA(p, q)$ as k tends to ∞ .
- Similarly, the variance of the forecast error converges to the unconditional variance.