Stochastic Processes

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Week 3

Examples of discrete random variables

Functions of discrete random variables

Expectation

Variance

Geometric

We say that X has the Geometric distribution (X is a Geometric r.v.) with parameter $p \in (0, 1)$ if X takes value $\{1, 2, ...\}$ and

$$P(X = k) = p q^{k-1}$$
 for $k = 1, 2, ...$

Observe that for $x \in \mathbb{R}$

$$(1-x)\sum_{k=0}^{n-1} x^{k} = \sum_{k=0}^{n-1} x^{k} - \sum_{k=1}^{n} x^{k}$$
$$= 1 + x + x^{2} + \dots + x^{n-1} - (x + x^{2} + \dots + x^{n}) = 1 - x^{n}$$

Then

$$\sum_{k=0}^{n-1} x^k = \frac{1-x^n}{1-x}$$

and

$$\sum_{k=0}^{\infty} x^{k} = \lim_{n \to \infty} \sum_{k=0}^{n-1} x^{k} = \lim_{n \to \infty} \frac{1-x^{n}}{1-x} = \begin{cases} \frac{1}{1-x} & |x| < 1\\ \infty & |x| \ge 1 \end{cases}$$

Hence we have also that the sum of the geometric probabilities is 1

$$\sum_{k=1}^{\infty} pq^{k-1} = p\sum_{k=0}^{\infty} q^k = prac{1}{1-q} = rac{p}{p} = 1$$

Exercise 1.17 Let $p_1, p_2, \ldots p_N$ non negative numbers such that $\sum_{i=1}^{N} p_i = 1$. Consider $\Omega = \{\omega_1, \ldots, \omega_N\}$ and let \mathcal{F} be the power set of Ω

$$Q(A) = \sum_{i:\omega_i \in A} p_i$$

is a probability measure

$$- \varphi(\mathcal{L}) = \lambda \quad \text{Imfact } \varphi(\mathcal{L}) = \overline{Z} \quad P_{1} = \overline{Z} \quad P_{1} = 1$$

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$$- \varphi(\mathcal{L}) = \gamma_{0} \quad \forall A \quad G_{\mathcal{F}} \quad \text{Imfact } \varphi(\mathcal{L}) = \overline{Z} \quad P_{1} = \gamma_{0}$$

$$\text{Since } P_{1} = \gamma_{0} \quad \forall i$$

$$- Let \quad A = \bigcup_{J \in \mathcal{J}} A_{J} \quad \text{where } A_{J} \cap A_{J} = \phi \quad J \neq J'$$

$$P(\mathcal{L}) = P(\bigcup_{J \in \mathcal{J}} A_{J}) = \sum_{i: w, \in \mathcal{J}} P_{i} = \sum_{J \in \mathcal{J}} P_{i} = \sum_{j \in \mathcal{J}} P_{i}$$

$$P(\mathcal{L}) = P(\mathcal{L}) \quad \varphi_{j} = \sum_{i: w, \in \mathcal{J}} P_{i} = \sum_{j \in \mathcal{J}} P_{i} = \sum_{j$$

Example 2.18, A coin is tossed *n* times. $|\Omega| = 2^n$, sngle outcomes are

$$\underbrace{TTHT\cdots TTH}_{n \ coins}$$

• Let $h(\omega)$ be the number of heads and $t(\omega)$ be the number of tails, $t(\omega) = n - h(\omega)$. Consider $p \in (0, 1)$ and q = 1 - p. Let \mathcal{F} be the power set

Take

$$p(\omega) = p^{h(\omega)}q^{t(\omega)}$$

and

$$P(A) = \sum_{\omega \in A} p(\omega)$$

Note that

$$\sum_{\omega \in \Omega} p(\omega) = \sum_{k=0}^{n} \sum_{\omega:h(\omega)=k} p(\omega) = \sum_{k=0}^{n} \sum_{\omega:h(\omega)=k} p^{h(\omega)} q^{t(\omega)}$$
$$= \sum_{k=0}^{n} \sum_{\omega:h(\omega)=k} p^{k} q^{n-k} = \sum_{k=0}^{n} p^{k} q^{n-k} \sum_{\omega:h(\omega)=k} 1$$
$$= \sum_{k=0}^{n} p^{k} q^{n-k} {n \choose k} = 1$$

Then P is a probability (by ex.1.17)

Let

$$X_i(\omega) = \begin{cases} 1 & \text{if the } i\text{th entry in } \omega \text{ is } H \\ 0 & \text{if the } i\text{th entry in } \omega \text{ is } T \end{cases}$$

$$P(X_{i} = 0) = P(\omega \in \Omega : \omega_{i} = T) = \sum_{\omega:\omega_{i}=T} p(\omega) = \sum_{\omega:\omega_{i}=T} p^{h(\omega)}q^{t(\omega)}$$
$$= \sum_{h=0}^{n-1} \sum_{\omega:h(\omega)=h,\omega_{i}=T} p^{h(\omega)}q^{t(\omega)} = \sum_{h=0}^{n-1} \sum_{\omega:h(\omega)=h,\omega_{i}=T} p^{h}q^{n-h}$$
$$= \sum_{h=0}^{n-1} p^{h}q^{n-h} \sum_{\omega:h(\omega)=h,\omega_{i}=T} 1 = \sum_{h=0}^{n-1} p^{h}q^{n-h} \binom{n-1}{h}$$
$$= \mathbf{q} \underbrace{\sum_{h=0}^{n-1} \binom{n-1}{h} p^{h}q^{n-1-h}}_{1} = \mathbf{q}$$

Hence, each X_i takes the values 0 and 1 and $P(X_i = 0) = q$ and $P(X_i = 1) = p$, that is

$$X_i \sim Bernoulli(p)$$
 $i = 1, \ldots n$

Example 2.18 (a coin is to need a times)

$$\Omega = \begin{cases}
HHHH...H., THH...H., ..., TTT...T \\
W= (W_{4}W_{2}...W_{n}) \\
W= (W_{4}$$

Remember that
$$P(A| = \overline{Z} P(\omega) = \overline{Z} P(\omega) = U(A)$$

$$P(S_{m}=K) = P(\omega \in \Omega : S_{m}(\omega) = K) = P(\omega \in \Omega : h(\omega) = K) =$$

$$= \sum_{\omega:h(\omega)=K} p^{k} (i-p)^{+(\omega)} = \sum_{\omega:h(\omega)=K} p^{k} (i-p)^{-k} = \binom{n}{k} p^{k} q^{n-k}$$

Poisson distribution as a limit of Binomial

Suppose that m is very large, p is very small and L=M.p does not go to zero or to infinity

 $P\left(S_{n}=K\right) = \binom{n}{k} p^{k} q^{n-k} = \frac{n!}{(n-k)! k!} p^{k} q^{n-k}$ $\sim \frac{n^{k}}{k!} \left(\frac{\lambda}{m}\right)^{k} \left(1-\frac{\lambda}{m}\right)^{n} \left(1-\frac{\lambda}{m}\right)^{-k} \sim \frac{\lambda^{k}}{k!} e^{\lambda}$ $\sim \frac{\lambda^{k}}{k!} re^{\lambda} re^{\lambda}$

Example 2.21 Geometric distribution

Consider the following experiment: we toss a coin until the first head turns up and then we stop

$$\Omega = \{H, TH, TTH, TTTH, \cdots\} \bigcup \{TTTT \cdots \}$$

Let $\omega = T^k H$ be the outcome with k heads and assume that

$$p(\omega)=q^k p \quad p\in (0,1) \quad q=1-p$$

For the outcome $\omega = T^{\infty}$ without heads we take

$$p(\omega)=\left\{egin{array}{cc} 1 & p=0\ 0 & p>0 \end{array}
ight.$$

Note that if p > 0

$$\sum_{\omega \in \Omega} p(\omega) = \sum_{k=0}^{\infty} p(T^{k}H) + p(T^{\infty}) = \sum_{k=0}^{\infty} pq^{k} + 0 = p\frac{1}{1-q} = 1$$

If p=0 we have also $\sum_{\omega\in\Omega}p(\omega)=0+0+0+\dots+1=1$

Consider the application

$$Y:\Omega
ightarrow \mathbb{R}$$

where

$$Y(\omega) = Y(T^k H) = k + 1$$

Then

 $Y \in \{1,2,\ldots\}$

and

$$P(Y = k) = P(T^{k-1}H) = q^{k-1}p$$
 $k = 1, 2, ...$

Hence Y has geometric distribution with parameter p

Exercise 2.24. Show that if Y is Geometric p, $P(Y > k) = (1 - p)^k$

Functions of discrete random variables

Let X be a discrete random variable on (Ω, \mathcal{F}, P) and let $g : \mathbb{R} \to \mathbb{R}$. Then Y = g(X) defined as

$$Y(\omega) = g(X(\omega))$$

is a random variable on (Ω, \mathcal{F}, P) Some examples

- if g(x) = ax + b then g(X) = aX + b
- if $g(x) = cx^2$ then $g(X) = cX^2$
- if $g(x) = \exp(x)$ then $g(X) = \exp(X)$

If
$$\gamma = g(X)$$
 the probability mans function of γ is
 $P_{\gamma}(y) = \sum_{x \in g^{-1}(y)} P(X=X)$
where $g^{-1}(y) = \{x:g(x) = y\}$. Infact
 $\gamma = y = \{g(X) = y\} = \{X \in \{x:g(x) = y\}\} = \{X \in g^{-1}(y)\}$
Then $P_{\gamma}(y) = P(g(X) = y\} = P(X \in \{x:g(x) = y\}) =$
 $= \sum_{x:g(X) = y} P(X=x) = \sum_{x \in g^{-1}(y)} P(X=x)$

• If
$$Y = aX + b$$
 then
 $P(Y = y) = P(aX + b = y) = P(X = (y - b)/a)$
• If $Y = X^2$ for $y > 0$
 $P(Y = y) = P(X^2 = y) = P(X = -\sqrt{y}) + P(X = \sqrt{y})$
Moreover for $y = 0$

$$P(Y = 0) = P(X^2 = 0) = P(X = 0)$$

and for y < 0

$$P(Y=y)=0$$

Expectation

Definition. If X is a discrete r.v. the **expectation** (expected value, mean) of X is denoted with E(X) and defined by

$$E(X) = \sum_{x \in Im(X)} x P(X = x)$$

whenever this sum converges absolutely in that $\sum_x |x \mathcal{P}(X=x)| < \infty$

Expecation for some important discrete random variables X

- Bernoulli(p), E(X) = p
- Binomial(n,p,) E(X) = np
- Poisson(λ), $E(X) = \lambda$
- Geometric(p), *E*(*X*) = 1/*p*
- Discrete uniform on $\{0, \dots n\}$, E(X) = n/2

 $X \sim \text{Bernoulli}(p)$

$$E(X) = \sum_{x} x P(X = x) = 0 \cdot (1 - p) + 1 \cdot p = p$$

 $X \sim \mathsf{Poisson}(\lambda)$

$$E(X) = \sum_{x} x P(X = x) = \sum_{x=0}^{\infty} x \frac{e^{-\lambda} \lambda^{x}}{x!}$$
$$= \sum_{x=1}^{\infty} x \frac{e^{-\lambda} \lambda^{x}}{x!} = \lambda \sum_{x=1}^{\infty} \frac{e^{-\lambda} \lambda^{(x-1)}}{(x-1)!}$$
$$= \lambda \sum_{y=0}^{\infty} \frac{e^{-\lambda} \lambda^{y}}{y!} = \lambda$$

THEOREM: If X is a discrete vandom variable and $g: R \rightarrow R$ then $E(g(X)) = \sum g(X) P(X=x)$ whenever this sum converges absolutely.



Proof
Let I the image of X. The image of Y is
$$g(I)$$

 $I = \sum_{X \in I} g(X)$
 $E(Y) = \sum_{X \in I} g(Y=y) = \sum_{Y \in g(I)} g(X) = y$
 $g \in g(I)$
 $F(X=x) = \sum_{X \in g(X)=y} g(X) = y$
 $g \in g(I) = \sum_{X \in g(X)=y} g \in g(X) = y$
 $g \in g(I) = \sum_{X \in g(X)=y} g \in g(X) = y$
 $f(X=x) = \sum_{X \in I} g(X) = y$

THEOREM

Let X be a discrete random variable. If $P(X \ge 0) = 1$ (X takes non negative values) and E(X) = 0, then P(X=0) = 1

PROOF

Infact
$$0 = \sum x P(X=x) = 0 \cdot P(X=o) + \sum x \cdot P(X=x)$$

Hence $\sum x P(X=x) = 0$ which is possible only
if $P(X=x) = 0 + x > 0$. There fore $P(X=o) = 1$

Exercise

Suppose that X is a discrete vandom variable with
mean
$$p = E(X)$$
. Consider $Y = aX + b$. Find $E(Y)$
 $Y = g(X)'$ where $g = ax + b$
 $E(Y) = \sum_{x} g(x) P(X = x) = \sum_{x} (ax + b) P(X = x) = \sum_{x} (ax + b) P(X = x) = \sum_{x} (ax + b) P(X = x) = \sum_{x} P(X = x) + \sum_{x} b P(X = x) = a \cdot \sum_{x} P(X = x) + b = ap + b$

Variance

Definition The variance var(X) of a discrete random variable X is defined by

$$\operatorname{var}(X) = E([X - E(X)]^2)$$

Consider $g(x) = (x - \mu)^2$ where $\mu = E(X)$ and remember that

$$E(g(X)) = \sum_{x \in Im(X)} g(x)P(X = x)$$

Then

$$\operatorname{var}(X) = \sum_{x \in \operatorname{Im}(X)} (x - \mu)^2 P(X = x)$$

Note that $var(X) = E(X^2) - E(X)^2$

$$\operatorname{var}(X) = \sum_{x} (x - \mu)^2 P(X = x)$$

= $\sum_{x} (x^2 - 2x\mu + \mu^2) P(X = x)$
= $\sum_{x} x^2 P(X = x) - 2\mu \sum_{x} x P(X = x) + \mu^2 \sum_{x} P(X = x)$
= $E(X^2) - 2\mu\mu + \mu^2 = E(X^2) - \mu^2 = E(X^2) - E(X)^2$

Consider
$$Y=aX$$
, then $var(Y) = a^{2}var(X)$
 $Julact E(Y) = E(aX) = aE(X)$ and
 $var(Y) = E(Y^{2}) - E(Y)^{2} = E(a^{2}X^{2}) - (aE(X))^{2}$
 $= a^{2}E(X^{2}) - a^{2}E(X)^{2} =$
 $= a^{2}(E(X^{2}) - E(X)^{2}) = a^{2}var(X)$

Exercise: Consider Y= aX+5, show that war (Y/= a var (X)

Variance for some important discrete random variables X

- Bernoulli(p), var(X) = pq
- Binomial(n,p,) var(X) = npq
- Poisson(λ), var(X) = λ
- Geometric(p), $var(X) = q/p^2$